CHILD SEXUAL ABUSE AND YOUTH SUICIDE: A REVIEW OF THE EVIDENCE WITH IMPLICATIONS FOR FUTURE RESEARCH

Anne E. Rhodes, Jennifer Bethell, and Lil Tonmyr

Abstract: Studies of suicide and non-fatal suicide-related behaviours demonstrate a gender paradox: Suicide rates are typically higher in males than females, whereas the opposite is true for non-fatal suicide-related behaviours. However, the reasons for these differences are unclear. Among the potential explanations, particularly in youth, is the effect of child maltreatment. A previous review suggested that while child sexual abuse may be more common in girls, the negative effect may be more potent for boys (with respect to suicide attempts). However, as their risk/protective factors may not always overlap, it is unclear whether this pattern seen for suicide attempts extends to suicide. The current study reviewed the evidence for potential sex differences in the association between child sexual abuse and suicide, identified methodological challenges to such studies, and discussed implications for future theoretical formulations/testing in research. The findings confirmed that the association between child sexual abuse and youth suicide remains unclear and the potential sex differences in the association remain largely unaddressed. Further, a test of the association between child sexual abuse and suicide would be best pursued in very large, population-based studies (with standardized measures of child sexual abuse) later linked to mortality data.

Keywords: suicide; self-injurious behavior; suicide, attempted; child abuse, sexual; child; adolescent

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Anne E. Rhodes, Ph.D. (the corresponding author) is an Associate Professor in Department of Psychiatry and the Dalla Lana School of Public Health, University of Toronto and a Research Scientist at the Suicide Studies Research Unit, St. Michael's Hospital, 30 Bond Street, Toronto, Ontario, M5B 1W8, Canada. E-mail: rhodesa@smh.ca

Jennifer Bethell, Ph.D. is a recent graduate of the doctoral program at the Dalla Lana School of Public Health, University of Toronto and works with the Suicide Studies Research Unit, St. Michael’s Hospital, Toronto, Canada. E-mail: bethellj@shm.ca

Lil Tonmyr, MSW, Ph.D. is a senior researcher with the Injury and Child Maltreatment Section, Centre for Chronic Disease Prevention and Control, Public Health Agency (PHAC) of Canada, Ottawa, Canada. E-mail: Lil.Tonmyr@phac-aspc.gc.ca
The gender paradox of suicide and non-fatal suicide-related behaviours

In most regions of the world, the suicide rate in males is at least double that of females and the sex differential is most apparent at older ages (Krug, Dahlberg, Mercy, Zwi, & Lozano, 2002). Conversely, non-fatal suicide-related behaviours (SRB) self-inflicted injury or poisoning tend to be more common in females and the sex differential lessens with age (Hawton & Harriss, 2008). However, the reasons for these sex differences, first evident in youth, have not been thoroughly explained. As suicide rates begin to increase during the transition from childhood to adolescence, making it the second leading cause of death in 10- to 24-year-olds worldwide (World Health Organization, 2010), studies of this age group are crucial.

The association between child maltreatment and suicide

Among the potential explanations for the sex differences in SRB rates, particularly in youth, is the effect of child maltreatment. Although this area of research has been somewhat limited with respect to deciphering the unique contribution of specific forms of abuse (Joiner et al., 2007), recent evidence suggests child sexual abuse (CSA) may be a particularly potent risk factor. Bruffaerts et al. (2010) used the World Mental Health Survey (55,299 respondents aged 18 years and over in 21 countries) to study the association between nine childhood adversities (physical abuse, sexual abuse, neglect, parental death, parent divorce, other parental loss, family violence, physical illness, and financial adversity before age of 18) and lifetime suicide ideation and attempts. The study showed CSA was the strongest predictor for both ideation and attempts, in bivariate and multivariate analyses adjusting for other adversities, as well as demographic and parental psychopathology variables. Similarly, CSA has been found to predict suicide attempts among suicidal ideators, independent of other factors (Brezo et al., 2007). Further, although other studies have demonstrated the association between other forms of maltreatment and non-fatal SRB in youth, they have not always controlled for CSA (Mironova et al., 2011), even though different forms of abuse may often present together and thus bias results and limit the studies’ abilities to test theories linking the abuse to outcomes (Joiner et al., 2007).

With regard to potential sex differences in the association between CSA and non-fatal SRB specifically in youth, in a previous systematic review (Rhodes et al., 2011), the relationship was found stronger in boys, and specifically for suicide attempts. Seven population-based studies provided sex-specific unadjusted estimates of the association between CSA and suicide attempt(s) and all found a statistically significant positive association in both boys and girls, but the association was stronger in boys than girls.

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1 Suicide-related behaviours (SRB) encompass fatal and non-fatal acts, including suicides (deaths with at least some degree of suicidal intent), other self-inflicted deaths, and non-fatal self-inflicted injury or poisoning. Non-fatal SRB includes suicide attempts (where there is an intent to die from suicide) as well as acts carried out without this intent (Silverman, Berman, Sanddal, O'Carroll, & Joiner, 2007a; Silverman, Berman, Sanddal, O'Carroll, & Joiner, 2007b).
(odds ratios [ORs] from 4.5 to 30.8 and 2.2 to 5.1, respectively), including for multiple suicide attempts (ORs 11.2 and 5.0, respectively). Fewer studies provided adjusted estimates; however, these studies showed that other factors explained some, but not all, of the observed sex difference. This sex differential may reflect the nature and timing of the abuse. For example, perpetrators of abuse of boys are more likely to be same-sex (Edgardh & Ormstad, 2000) which may be especially traumatic and isolating for these victims, also making them less likely to disclose the abuse and seek help because of shame or fear of stigmatization. Other factors may include sex differences in whether the abuse involved the use of physical force; Joiner’s theory and analysis of suicidal behaviour have implicated exposure to violence in the context of abuse as key to later risk for suicidal behaviour (Joiner et al., 2007). In other words, while CSA may be more common in girls, the negative effect may somehow be stronger for boys (with respect to suicide attempts). Taken together, an implication for population attributable risk estimates may be that CSA gives rise to as many suicide attempts among boys exposed to CSA as it does among exposed girls (Rhodes et al., 2011).

Although a stronger association between CSA and suicide attempt(s) among boys compared to girls has already been reported, as their risk/protective factors may not always overlap, it is unclear whether this pattern seen for suicide attempt(s) extends to suicide. It is critical to clarify the validity of causal pathways to inform prevention efforts. Accordingly, this report will begin to assess sex differences in the association between CSA and youth suicide and discuss implications for future theoretical formulations/testing in research.

**The Present Investigation**

This study reviewed the empirical literature on the association between CSA and suicide in samples of youth, highlighting evidence of possible sex differences in the association, then identifying methodological challenges to such studies and discussing implications for future research.

**Method**

A search of published research was carried out. Studies eligible for inclusion met each of the following criteria: published between 1988 and 2011 (inclusive); reported empirical results on the estimate of CSA among suicides, and/or the association between CSA and suicide; included (but were not necessarily limited to) suicides in children, youth and young adults (up to age 25 years); and were English language. Electronic databases (including Medline, PsycINFO, and Social Services Abstracts) were searched, using subject headings “suicide” and “child abuse, sexual”, to identify relevant papers. Case reports, qualitative studies, reviews, and editorials were excluded, along with studies that examined non-fatal SRB and/or suicidal ideation (not suicide), child welfare involvement (not specified as CSA), or, suicide in CSA perpetrators (not victims). Additional papers were identified by searching references lists. Descriptive information
on the studies was abstracted and tabulated by one reviewer: type of study (psychological autopsy\(^2\) or cohort study); information about the population studied; study characteristics (whether a control group was included and whether analyses were sex-stratified); and estimates of CSA in youth suicides (including the association with suicide). This information was reviewed and checked for accuracy by a second reviewer. Formal quality assessment rules were not applied, given the lack of consensus and evaluation tools to assess the quality of observational studies (Sanderson, Tatt, & Higgins, 2007).

**Results**

The search yielded nine studies. Tables 1 and 2 provide information on the six psychological autopsy studies (Séguin, Renaud, Lesage, Robert, & Turecki, 2011; Moskos, Olson, Halbern, Keller, & Gray, 2005; Houston, Hawton, & Shepperd, 2001; Appleby, Cooper, Amos, & Faragher, 1999; Brent, Baugher, Bridge, Chen, & Chiapetta, 1999; Brent et al., 1994) and three cohort studies (Cutajar et al., 2010; White & Widom, 2003; Plunkett et al., 2001) that reported estimates of CSA among suicides, and/or estimates of the association between CSA and suicide. However, of these nine studies, three (Séguin et al., 2011; Brent et al., 1999; White & Widom, 2003) did not distinguish sexual abuse from other types of abuse and maltreatment. Publication bias appears unlikely among the psychological autopsy studies, given that CSA was typically studied among an array of potential risk factors.

Overall, of these nine studies, only three reported sex-stratified results: one psychological autopsy study (Brent et al., 1999) and two cohort studies (White & Widom, 2003; Cutajar et al., 2010). In the psychological autopsy studies, where reported, estimates of the history of CSA among suicides were about 20% (Houston et al., 2001; Appleby et al., 1999), but ranged from less than 10% (Brent et al., 1994) to 33% (Moskos et al., 2005) (and no sex-specific estimates of CSA were provided). Moskos et al. (2005) found history of CSA ranged from about 5% to 33%, depending on the informant type (CSA was reported most often by parents). Information on the perpetrator was reported only by Houston et al. (2001), who found that of the five cases of CSA, the perpetrator was a parent in one, and for the remainder, perpetrators were friends, family friends, and strangers. Estimates of the history of CSA among suicides in the cohort studies could not be derived.

Very few psychological autopsy studies provided estimates of the association between CSA and suicide, either because there was no control group (Moskos et al., 2005; Houston et al., 2001) or because there were empty cells for the control group (Appleby et al., 1999). Of the two studies that did provide estimates of the association (Séguin et al., 2011; Brent et al., 1999), neither distinguished CSA from other types of abuse and maltreatment and both were hampered by very small numbers (especially in the control group), yielding very wide confidence intervals for effect estimates.

\(^2\) A systematic assessment of decedents’ circumstances before suicide, using interviews with one or more proxy respondents (i.e., informants) (Conner et al., 2012)
Estimates of the association between CSA and suicide from cohort studies varied; White and Widom (2003) showed neither a statistically significant association between maltreatment and all-cause or violent death (suicide or homicide) mortality, Plunkett et al. (2001) found the suicide rate among CSA victims was about 10 times the comparable national rate, and Cutajar et al. (2010) showed the relative risk of suicide, comparing CSA victims to controls, was 14.2 (95% CI: 5.0-40.6) in males and 40.4 (95% CI: 25.0-65.3) in females, however, subjects were followed for many years, some up until age 64.

Taken together, a history of CSA appears to be fairly common among youth who die by suicide (reported for roughly one in five). Cohort studies of CSA victims estimate suicide risk may be at least tenfold higher compared to the general population. However, it remains unclear whether the stronger association between CSA and suicide attempt(s) observed for boys compared to girls (Rhodes et al., 2011) extends to youth suicide.
**Table 1. Psychological autopsy studies of suicide that reported on sexual abuse**

<table>
<thead>
<tr>
<th>Author, year</th>
<th>Population (suicide cases)</th>
<th>Study characteristics</th>
<th>Estimate of sexual abuse among suicides</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Location, Age (years)</td>
<td>Source, year of death</td>
</tr>
<tr>
<td>Séguin et al., 2011</td>
<td>Quebec, Canada, &lt;30</td>
<td>Consecutive suicides recorded by Quebec’s Coroner’s Office and the Montreal Central Morgue, year of death not stated</td>
<td>67 of 89 (75%)</td>
</tr>
<tr>
<td>Moskos et al., 2005</td>
<td>Utah, U.S.A., 13-21</td>
<td>151 consecutive suicides recorded by Utah Office of the Medical Examiner, 1996-98</td>
<td>51 of 108 families contacted (47%)</td>
</tr>
<tr>
<td>Houston et al., 2001</td>
<td>Oxford-area, England, 15-24</td>
<td>35 consecutive suicides, 26 undetermined deaths and 26 accidental deaths, recorded by</td>
<td>27 of 47 (57%)</td>
</tr>
<tr>
<td>Study</td>
<td>Location</td>
<td>Age</td>
<td>Type of Study</td>
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<tr>
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</tr>
<tr>
<td>Appleby et al., 1999</td>
<td>Manchester, England</td>
<td>&lt;35</td>
<td>Consecutive suicides and undetermined deaths</td>
</tr>
<tr>
<td>Brent et al., 1999</td>
<td>Western Pennsylvania, U.S.A.</td>
<td>13-19</td>
<td>Consecutive suicides, year of death not stated</td>
</tr>
<tr>
<td>Brent et al., 1994</td>
<td>Western Pennsylvania, U.S.A.</td>
<td>&lt;20</td>
<td>Consecutive suicides, 1986-90</td>
</tr>
<tr>
<td>U.S.A.</td>
<td></td>
<td></td>
<td>geographic cluster sampling, matched on age, sex, county of residence and SES</td>
</tr>
</tbody>
</table>

OR: Odds Ratio; SES: socioeconomic status; 95% CI: 95% Confidence Interval
Table 2. Cohort studies of child sexual abuse (CSA) and suicide

<table>
<thead>
<tr>
<th>Author, year</th>
<th>Population (sexual abuse cases)</th>
<th>Death data source</th>
<th>Study characteristics</th>
<th>Suicide deaths</th>
</tr>
</thead>
<tbody>
<tr>
<td>Cutajar et al., 2010</td>
<td>Victoria, Australia, &lt;17, 13-44 years</td>
<td>National Coroners Information System (NCIS) and Victorian Coronial Information Database (VCID)</td>
<td>Australian national population data, for 15-64 year-olds, for the year 2000</td>
<td>Eight deaths by suicide documented in the sexually abused cohort, two (0.4%) in males and six (0.3%) in females. RR suicide (CSA vs. control cohort): 14.2 (95% CI: 5.0-40.6) in males and 40.4 (95% CI: 25.0-65.3) in females.</td>
</tr>
<tr>
<td>White &amp; Widom, 2003</td>
<td>Metropolitan area of Midwest, U.S.A., &lt;11, 23-27 years</td>
<td>National Death Index</td>
<td>667 children matched on age, sex, race/ethnicity, family social class.</td>
<td>Yes, for all-cause mortality</td>
</tr>
</tbody>
</table>

RR violent death, i.e., homicide or suicide (abused vs. control group): 0.4 (95% CI: 0.1-1.9)
<table>
<thead>
<tr>
<th>Study</th>
<th>Location</th>
<th>Age</th>
<th>Gender</th>
<th>Design Description</th>
<th>Control Group</th>
<th>Suicide Rate</th>
</tr>
</thead>
<tbody>
<tr>
<td>Plunkett et al., 2001</td>
<td>Sydney, Australia</td>
<td>4-15</td>
<td>Female</td>
<td>183 children presenting to one of two specialist Child Protection Units in Sydney in 1998 and 1999 and confirmed as sexually abused.</td>
<td>No</td>
<td>Three deaths by suicide in the sexually abused cohort, none among control cohort. Suicide rate among abused: 179 per 100,000 person-years (comparable national suicide rate: 13.8-16.7 per 100,000 person-years)</td>
</tr>
</tbody>
</table>

RR: Relative Risk; 95% CI: 95% Confidence Interval
Discussion

This literature review demonstrated there are very few studies of the association between CSA and suicide. Accordingly, we have a poor understanding of the relationship and any potential sex differences. The few studies that have tested the association between CSA and suicide have been hampered by methodological challenges. The following section will outline these methodological challenges, by study type (as applicable).

Sampling

**Psychological autopsy studies.** Moskos et al. (2005) found that of the families invited to participate in the study, those who had been referred to child and protective services were less likely to participate. Given that response rates for psychological autopsy studies tended to be about 50% to 75%, this could have important implications. More specifically, this would suggest that CSA (and other forms of maltreatment) would be underestimated among youth suicides studied by psychological autopsy. It is also unclear whether this pattern of study participation also applies to controls and/or differs by sex.

**Cohort studies.** These studies included children and youth for whom abuse had been substantiated through medical, legal and/or social services. As such, they represented only the very small subset of abused children and youth who come to the attention of authorities (MacMillan, Jamieson, & Walsh, 2003). Further, as discussed by Cutajar et al. (2010), these children tend to come from more “disorganized and disadvantaged families”; among children exposed to CSA, those from lower income families may be more likely to come into contact with child protection services (MacMillan et al., 2003). Children in these families likely also face other factors, aside from abuse, that might put them at increased risk for suicide (thereby possibly explaining the observed positive association). It is also unclear whether sexually abused girls are more or less likely than sexually abused boys to have their abuse substantiated through these services. Given that boys may be less likely to discuss their abuse (Hebert, Tourigny, Cyr, McDuff, & Joly, 2009; Edgardh & Ormstad, 2000), it seems plausible that official records capture a smaller proportion of abused boys than abused girls; in fact, MacMillan et al. (2003) showed that among children who reported CSA, girls were more likely to report contact with child protection services than boys, but the effect was not statistically significant (odds ratio 1.70, 95% confidence interval 0.71-4.08).

Misclassification

**Psychological autopsy studies.** Methodological problems, including differences in measures and the number and type of informants used for each suicide, all of which vary across studies, in all likelihood affect the comparability and reliability of data from psychological autopsy studies (Pouliot & De Leo, 2006). In fact, Moskos et al. (2005)
found that the history of CSA depended on the type of informant; CSA was frequently reported by parents (33%), but much less often by others. CSA is highly secretive, so measurement may be particularly susceptible to misclassification either because the informant was unaware or deliberately withheld some information (Conner et al., 2012), for example, because most CSA is perpetrated by someone known to the victim, including family (Finkelhor, 1994). As such, the history of CSA is likely under-recorded among youth suicides (as well as controls), and assuming non-differential misclassification, estimates of the association between CSA and suicide are likely also underestimated (Kleinbaum, Kupper, & Morgenstern, 1982). Furthermore, given that boys appear less likely to disclose CSA (Edgardh & Ormstad, 2000), this misclassification may differ by sex.

**Cohort studies.** Although the substantiated CSA cases may be well defined in the cohort studies, control groups no doubt included some victims of CSA. For example, Cutajar et al. (2010) used national data as the control population; however, given that CSA cases were restricted to only those documented in the State of Victoria (population about 5 million), the control group would include not only the substantial proportion of CSA that goes undocumented, but also CSA victims documented in other jurisdictions. As such, the authors themselves pointed out that their study likely underestimated the association between CSA and suicide. Again, given boys may be less likely to disclose their CSA, this misclassification may differ by sex.

**Sample Size and Statistical Power**

**Psychological autopsy studies.** These studies typically do not accrue large samples because of the time and resources required to do so. As such, studying relatively rare exposures, such as CSA, may be problematic for analysis and/or yield imprecise effect estimates. For example, several studies found that none of the control group had a history of CSA and thus, in the analysis, the empty cells would have caused bias in odds ratios estimates (Agresti, 1996). This issue also restricts opportunity for sex-stratified analyses.

**Cohort studies.** Even assuming suicide risk is elevated among children and youth exposed to CSA, fortunately suicide is still a rare outcome and therefore, requires long follow-up and/or large samples exposed and unexposed to CSA in order to accumulate events. For example, Cutajar et al. (2010) followed a cohort of 2,759 individuals with CSA for up to 44 years, and recorded eight suicides (< 0.5% of the cohort).

**Theoretical Formulations**

A better understanding of the complex association between CSA and suicide will require identifying and analyzing related factors (linked to both CSA and suicidal ideation, plans, and behaviour) (Kalucy, 2010), including mental health problems (Green, 1993), re-victimization (Balsam, Lehavot, & Beadnell, 2011) and other risk/protective exposures during childhood (Nelson et al., 2002). As such, future studies will need to
carefully consider the roles of other variables in the association between CSA and suicide (e.g., confounder or mediator), including how to analyse them in the context of a wider hypothesized causal model (Vittinghoff, Glidden, Shiboski, & McCulloch, 2005). For example, one study showed the association between CSA and suicide ideation and attempts was largely mediated by mental health problems and stressful life events (Fergusson, Woodward, & Horwood, 2000). Conversely, Molnar, Berkman, and Buka (2001) found the association was only partially explained by psychiatric disorders and other childhood adversities (e.g., other forms of parental abuse). With regard to any sex differences in the association between CSA and suicide, Rhodes et al. (2011) identified possible pathways for further investigation, suggesting differential disclosure and help-seeking behaviour and recommending qualitative research, mixed methods, and more focused hypothesis testing in large, prospective observational studies.

**Suicides in Adulthood**

It should be noted that the cohort studies listed in the previous section extended follow-up into adulthood and therefore, may not have been studying strictly youth suicide as the outcome. In fact, Cutajar et al. (2010) found most of the suicides occurred in adulthood (mean age: 31). Accordingly, it is unclear if these relative risk estimates apply specifically to youth. Further, although we have previously reported that the association between CSA and suicide attempts may be stronger in boys than in girls, this finding may not extend to adults; results from large surveys in mainly adults have reported differing findings with regard to sex differences in the association between sexual abuse and suicide attempts (Bebbington et al., 2009; Molnar et al., 2001), suggesting substantive differences between CSA and lifetime sexual abuse and/or methodological issues including recall, disclosure, or selection bias (e.g., because of premature mortality in males). Taken together, regardless of study type, these findings reinforce the importance of carefully considering and specifying the age at which the abuse occurs and (suicide and non-fatal SRB) outcomes are studied, including the need for stratification by age as well as sex.

**Suicides in Childhood**

CSA appears to be associated with non-fatal SRB at a young age, perhaps even more strongly than later in life. Bruffaerts et al.’s analysis of World Mental Health survey data showed that history of CSA (before age 18) was associated with a tenfold increase in the odds of a suicide attempt between ages 4 and 12, but this association decreased for suicide attempts in adolescence and adulthood, with odds ratios of about 6 and 3, respectively (Bruffaerts et al., 2010). Similarly, analysis of the Netherlands Mental Health Survey and Incidence Study, with just over 7,000 adults, did not find an association between CSA and incident suicide attempts, however, additional analyses suggested this was because CSA likely results in onset of suicide attempts at a relatively young age (Enns et al., 2006). Nevertheless, there are considerable concerns in assigning a verdict of suicide to children, such as whether they communicate or understand suicidal
intent (Crepeau-Hobson, 2010). Consequently, by excluding suicides that occur among the very young, studies may not capture the full extent of the association between CSA and suicide.

**Conclusion**

A review of the literature confirmed that the association between CSA and youth suicide remains unclear and the potential sex difference in the association remains largely unaddressed. Studies of the association between CSA and suicide face major methodological challenges, as discussed here, underscoring inherent difficulties in addressing these questions. Some of these challenges introduce opportunities for bias, and all require careful consideration when planning for such a study, including study design and population, data sources, sample size, and length of follow-up (if applicable).

Taken together, these findings suggest that the theory regarding sex differences in the association between CSA and suicide needs further investigation. Ideally, theories could be tested in large, population-based cohorts with standardized measures of CSA, later linked to mortality data. Two such examples are population-based surveys, and, population-based child welfare registries or surveys, similar to Cutajar et al. (2010). Population-based surveys may be most promising because surveys in youth have already demonstrated sex differences in the association between CSA and suicide attempts and can capture CSA that has not been formally disclosed. Given the likely sex differential in disclosure of CSA (girls disclose more than boys), it is possible that studies based on child welfare registries or child welfare surveys would not find sex differences in the association between CSA and suicide, for example, if shame and isolation from non-disclosure were key in the causal pathway to suicide. However, to date, the surveys of youth that have addressed CSA and non-fatal SRB have been anonymous, in school-based settings (Rhodes et al., 2011) or collected data retrospectively from adults in order to avoid ethical issues of mandatory reporting of child abuse (Fergusson et al., 2000). The designs in the former may have precluded later linkage to mortality data, and the latter may introduce opportunities for bias (e.g., recall, disclosure, or selection). Ultimately, the need to test theoretical formulations to explain the association between CSA and suicide, and any potential sex differences therein, must be approached with an appreciation of the challenges in conducting the research and interpreting the study results.
References


